

# Combined Stress Scenarios for Market and Credit Risks

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Generation of Random Stress Scenarios

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A Company's Value in Respect to the Merton Model/Credit Metrics  
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# 1. Multi-firm Merton model

## 1.1 The asset value model

Portfolio of  $n$  firms with time horizon  $T$ .

- $S_t^{(i)}$ : equity value of firm  $i \in \{1, \dots, n\}$  at time  $t \in [0, T]$ ,
- $F_t^{(i)}$ : value of the liabilities of firm  $i$  at time  $t \in [0, T]$ ,
- $V_t^{(i)}$ : value of firm  $i$ 's assets at time  $t \in [0, T]$ .

## Assumptions

- $T_i$ : time to maturity of liabilities of firm  $i$ .
- $H^{(i)}$ : face value of liabilities of firm  $i$ .
- $S_{T_i}^{(i)} := \max(V_{T_i}^{(i)} - H^{(i)}, 0)$  is the amount that the equity holders would receive at maturity  $T_i$  in case of liquidation of firm  $i$ .

- $F_{T_i}^{(i)} := \min(V_{T_i}^{(i)}, H^{(i)}) = H^{(i)} - \max(H^{(i)} - V_{T_i}^{(i)}, 0)$  value of the bond at  $T_i$ . Default if  $V_{T_i}^{(i)} < H^{(i)}$ .
- $V_t^{(i)} = S_t^{(i)} + F_t^{(i)}$  total value of firm  $i$  at time  $t \in [0, T]$ .
- Dynamics of asset value processes:

$$dV_t^{(i)} = \mu_{V_i} V_t^{(i)} dt + \sigma_{V_i} V_t^{(i)} dW_t^{(i)} \quad (1)$$

with  $\mu_{V_i}$  instantaneous expected rate of return to the firm per time unit,

$\sigma_{V_i}^2$  instantaneous variance of the logarithmic firm value per time unit,

$W_t = (W_t^{(1)}, \dots, W_t^{(n)})$   $n$ -dimensional Brownian motion.

- Complete market. Under unique martingale measure  $\mathbb{Q}$

$$dV_t^{(i)} = r V_t^{(i)} dt + \sigma_{V_i} V_t^{(i)} dW_t^{*,(i)}, \quad i = 1, \dots, n, \quad (2)$$

where  $r$  risk-less interest rate and

$W_t^* = (W_t^{*,(1)}, \dots, W_t^{*,(n)})$   $n$ -dimensional Brownian motion.

Prices  $S_t^{(i)}$  and  $F_t^{(i)}$  for equities and bonds

$$S_t^{(i)} = e^{r(t-T_i)} \mathbb{E}_{\mathbb{Q}}[S_{T_i}^{(i)} | \mathcal{F}_t] \quad (3)$$

$$F_t^{(i)} = e^{r(t-T_i)} \mathbb{E}_{\mathbb{Q}}[F_{T_i}^{(i)} | \mathcal{F}_t] \quad (4)$$

Yield spread:

$$Y_0^{(i)} \stackrel{\text{def}}{=} -\frac{1}{T_i} \log \frac{F_0^{(i)}}{H^{(i)}} - r.$$

## 1.2 The joint assets value distribution\*

Substituting  $\sum_{j=1}^M \sigma_{j,V_i} d\bar{W}_t^{(j)}$  for  $dW_t^{(i)}$  into (1) yields:

$$V_t^{(i)} = V_0^{(i)} \exp \left( \mu_{V_i} t - \frac{(\sigma_{V_i})^2}{2} t + \sum_{j=1}^M \sigma_{j,V_i} \bar{W}_t^{(j)} \right), \quad i = 1, \dots, n, \quad (5)$$

where  $\sigma_{V_i}$  satisfies  $(\sigma_{V_i})^2 := \sum_{j=1}^M (\sigma_{j,V_i})^2$ ,  
if  $(\bar{W}^{(j)})_{j=1,\dots,M}$  independent Brownian motions.

Note  $\underline{Z} \sim \mathcal{N}(\underline{0}, \underline{\Sigma})$  for  $Z_i := \frac{\sum_{j=1}^M \sigma_{j,V_i} \bar{W}_T^{(j)}}{\sigma_{V_i} \sqrt{T}}$  with

$$[\underline{\Sigma}]_{ij} = \frac{\sum_{s=1}^M \sigma_{s,V_i} \sigma_{s,V_j}}{\sigma_{V_i} \sigma_{V_j}}. \quad (6)$$

In particular  $\text{var}(Z_i) = 1$ .

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\*See M. Nyfeler's Diploma thesis

Estimation of  $\Sigma$ :

$$\text{Corr}\left(\log\left(\frac{V_T^{(i)}}{V_0^{(i)}}\right), \log\left(\frac{V_T^{(j)}}{V_0^{(j)}}\right)\right) = [\Sigma]_{ij}.$$

Number of parameters to estimate:  $2n + \frac{(n-1)n}{2}$ .

Reduction of dimension by factor model:

$\Sigma = \mathbf{A}\mathbf{A}^T + \mathbf{D}$  with

$$[\mathbf{A}]_{ij} = \frac{\sigma_{j,V_i}}{\sigma_{V_i}}, \quad [\mathbf{D}]_{ij} = \begin{cases} \left(\frac{\sigma_{K+i,V_i}}{\sigma_{V_i}}\right)^2 & \text{if } i = j, \\ 0 & \text{if } i \neq j. \end{cases}$$

Conclusion:  $Z$  can be represented by the factor model

$$Z_i \stackrel{d}{=} \sum_{j=1}^K [\mathbf{A}]_{ij} R_j + \mathcal{E}_i \quad i = 1, \dots, n \quad (7)$$

where

$$\begin{aligned} M &= K + n \\ \{(R_j)_j, (\mathcal{E}_i)_i\} &\text{ independent} \\ R_j &\sim \mathcal{N}(0, 1), \quad j = 1, \dots, K \\ \mathcal{E}_i &\sim \mathcal{N}(0, [\mathbf{D}]_{ii}), \quad i = 1, \dots, n. \end{aligned} \quad (8)$$



## 1.3 The loan-loss distribution

- Intention: determining distribution of *loan portfolio value*

$$\sum_{i=1}^n F_T^{(i)} \quad (9)$$

Portfolio loss: difference of benchmark value and loan portfolio value

- Analytical approximation: under simplifying assumptions for large portfolio
- Monte Carlo simulation by using a factor model

## 1.4 Stress Scenarios

(i) Provide data for every firm  $i \in \{1, \dots, n\}$ :

$r$ : risk-less interest rate (from market)

$T_i$ : average maturity of outstanding debt  
(book value)

$H^{(i)}$ : face value of outstanding debt (book value)

$S_0^{(i)}$ : equity value (from market)

$\mathbb{P}(V_T^{(i)} < H^{(i)})$ : physical default probability (by rating)

$Y_0^{(i)}$ : yield spread (by rating from market)

$V_0^{(i)}, \sigma_{V_i}$ : assets value and assets volatility  
(by numerical computation from Merton model)

(ii) Value of outstanding debt for firm  $i$  at time  $T$ :

$$F_T^{(i)} = \begin{cases} e^{r(T-T_i)} H^{(i)} \mathbb{I}_{\{V_{T_i}^{(i)} \geq H^{(i)}\}} & , \quad T \geq T_i \\ e^{r(T-T_i)} H^{(i)} \left( 1 - N \left( \frac{\log \frac{H^{(i)}}{v} + (\sigma_{V_i}^2 / 2 - r)(T_i - T)}{\sigma_{V_i} \sqrt{T_i - T}} \right) \right) \Big|_{v=V_T^{(i)}} & , \quad T < T_i \end{cases} \quad (10)$$

(iii) Value of bond portfolio:  $\sum_{i=1}^n F_T^{(i)}$

Monte Carlo simulation with (10).

Simulate  $V_{T_i}^{(i)}$  if  $T \geq T_i$ ,  $V_T^{(i)}$  if  $T < T_i$ .

(iv) Shock scenarios:

Equity, interest:  $S_0^{i,\text{new}} = (S_0^{i,\text{old}} + \Delta S) e^{-\Delta r T}$

Volatility:  $\sigma_{V_i}^{\text{new}} = \sigma_{V_i}^{\text{old}} (1 + \Delta \sigma)$

## 2 Integrating foreign exchange risk into the credit default model

### 2.1 Model based on three risky assets

#### Assumptions

- Constant domestic and foreign risk-free interest rates.
- Stochastic processes as geometric Brownian motions.
- Dependence by driving multidimensional Brownian motion.

## Approximate Black–Scholes valuation formula for basket options

Basket of tradable assets:

$$w_1 A_t^{(1)} + w_2 A_t^{(2)}, \quad w_1 + w_2 = 1, \quad w_1, w_2 \geq 0$$

Strike price:  $C$

Call option:

$$C_T = \left( w_1 A_t^{(1)} + w_2 A_t^{(2)} - C \right)^+ \quad (11)$$

In general:  $w_1 A_t^{(1)} + w_2 A_t^{(2)}$  not log-normal

⇒

Approximation by geometric mean  $(\hat{A}_t^{(1)})^{\hat{w}_1} (\hat{A}_t^{(2)})^{\hat{w}_2}$  of appropriately transformed  $\hat{A}_t^{(i)}$ ,  $\hat{w}_i$  (Gentle (1993))

⇒

Black–Scholes formula applicable, if market complete

## Notation for credit default model

$X_t$ : Exchange rate process representing the domestic price at time  $t$  of one unit of the foreign currency.

$V_t^d$ : Value of the domestic part of the firm's assets at time  $t$ .

$V_t^f$ : Value of the foreign part of the firm's assets at time  $t$  in foreign currency.

$r_d$ : Domestic risk-free interest rate.

$r_f$ : Foreign risk-free interest rate.

$B_t^d$ : Value of risk-free domestic bond at time  $t$ .

$B_t^f$ : Value of risk-free foreign bond at time  $t$ .

Dynamics under physical probability  $\mathbb{P}$ :

$$\begin{aligned}dX_t &= X_t[\mu_X dt + \sigma_1 \cdot dU_t] \\dV_t^d &= V_t^d[\mu_{V^d} dt + \sigma_2 \cdot dU_t] \\dV_t^f &= V_t^f[\mu_{V^f} dt + \sigma_3 \cdot dU_t] \\dB_t^d &= r_d B_t^d dt \\dB_t^f &= r_f B_t^f dt\end{aligned}$$

where

$$\begin{aligned}U_t &= (U_t^1, U_t^2, U_t^3) \text{ is a standard BM,} \\ \sigma_i &= (\sigma_{i1}, \sigma_{i2}, \sigma_{i3}) \in \mathbb{R}^3, \quad i = 1, 2, 3, \\ \sigma_i \cdot dU_t &= \sum_{j=1}^3 \sigma_{ij} dU_t^j.\end{aligned}$$

Unique domestic martingale measure ( $\Rightarrow$  complete market):

Girsanov theorem:

probability  $\mathbb{Q}$  such that  $\frac{V^d}{B^d}$ ,  $\frac{XV^f}{B^d}$ ,  $\frac{XB^f}{B^d}$  are martingales.

## Modeling assets with foreign exchange risk in practice

“Total assets = domestic assets + foreign assets”

$$V_t = V_t^d + \bar{V}_t^f, \quad \bar{V}_t^f = X_t V_t^f$$

Decomposition of  $V_t$ :

$$V_0^d = (1 - \alpha) V_0, \quad \bar{V}_0^f = \alpha V_0 \quad (12)$$

Take determination coefficient of regression of stock returns versus exchange rate returns as  $\alpha$ .

Take correlation of stock returns and foreign currency returns as approximation for correlation of returns to  $V_t^d$  and  $\bar{V}_t^f$ .

Substitute  $V_t^d$  and  $\bar{V}_t^f$  for  $A_t^{(1)}$  and  $A_t^{(2)}$  in Black–Scholes formula for basket option.

Stress scenarios for foreign exchange risk:

$$\bar{V}_0^{f,\text{new}} = \bar{V}_0^{f,\text{old}} (1 + \Delta)$$



## Problems with three-assets approach

- No obvious decomposition of assets value into domestic and foreign parts
- Estimation from regression of stock returns versus exchange rate returns does not seem reliable
- Alternative: estimate weight  $\alpha$  from balance sheet. Impractical. Does not yield correlation of domestic and foreign returns.

## 2.2 Model based on two risky assets

Modeling directly foreign exchange shocks difficult

⇒

indirect modeling by damped volatility shocks

### Notation

$U_t$ : 2-dimensional Brownian motion  $(U_t^1, U_t^2)$ .

$X_t$ : Exchange rate process representing the domestic price at time  $t$  of one unit of the foreign currency.

$V_t$ : Value of the firm's assets at time  $t$  in domestic currency.

$r_d$ : Domestic risk-free interest rate.

$r_f$ : Foreign risk-free interest rate.

$B_t^d$ : Value of risk-free domestic bond at time  $t$ .

$B_t^f$ : Value of risk-free foreign bond at time  $t$ .

Dynamics under physical probability  $\mathbb{P}$ :

$$dX_t = X_t[\mu_X dt + \sigma_1 \cdot dU_t]$$

$$dV_t = V_t[\mu_V dt + \sigma_2 \cdot dU_t]$$

$$dB_t^d = r_d B_t^d dt$$

$$dB_t^f = r_f B_t^f dt$$

where

$$U_t = (U_t^1, U_t^2) \quad \text{standard BM}$$

$$\sigma_i = (\sigma_{i1}, \sigma_{i2}) \in \mathbb{R}^2, \quad i = 1, 2,$$

$$\sigma_i \cdot dU_t = \sigma_{i1} dU_t^1 + \sigma_{i2} dU_t^2.$$

Unique domestic martingale measure ( $\Rightarrow$  complete market):

Girsanov theorem:

probability  $\mathbb{Q}$  such that  $\frac{V^d}{B^d}$  and  $\frac{XB^f}{B^d}$  are martingales.

$\Rightarrow$  “classical” Merton formula for value of equity.

Alternative representation for  $X_t, V_t$ :

$$\begin{aligned}dX_t &= X_t[\mu_X dt + \sigma_X dW_t^1] \\dV_t &= V_t[\mu_V dt + \sigma_V dW_t^2]\end{aligned}$$

with

$$W_t^1 = \sigma_{11} U_t^1 + \sigma_{12} U_t^2, \quad W_t^2 = \sigma_{21} U_t^1 + \sigma_{22} U_t^2.$$

Here

$$\text{corr}(W_t^1, W_t^2) = \sigma_{11} \sigma_{21} + \sigma_{12} \sigma_{22} \stackrel{\text{def}}{=} \rho_{XV}$$

and

$$\sigma_X = \sqrt{\sigma_{11}^2 + \sigma_{12}^2}, \quad \sigma_V = \sqrt{\sigma_{21}^2 + \sigma_{22}^2}.$$

## Generating stress scenarios

Assumption:  $\sigma_{12} = \sigma_{21}$

Then

$$\begin{aligned}\sigma_{11}^2 + \sigma_{12}^2 &= \sigma_X^2, \\ \sigma_{12}^2 + \sigma_{22}^2 &= \sigma_V^2\end{aligned}$$

Interpret foreign exchange shock as shock on  $\sigma_X$ :

$$\sigma_X^{\text{new}} = \sigma_X^{\text{old}}(1 + \Delta\sigma_X).$$

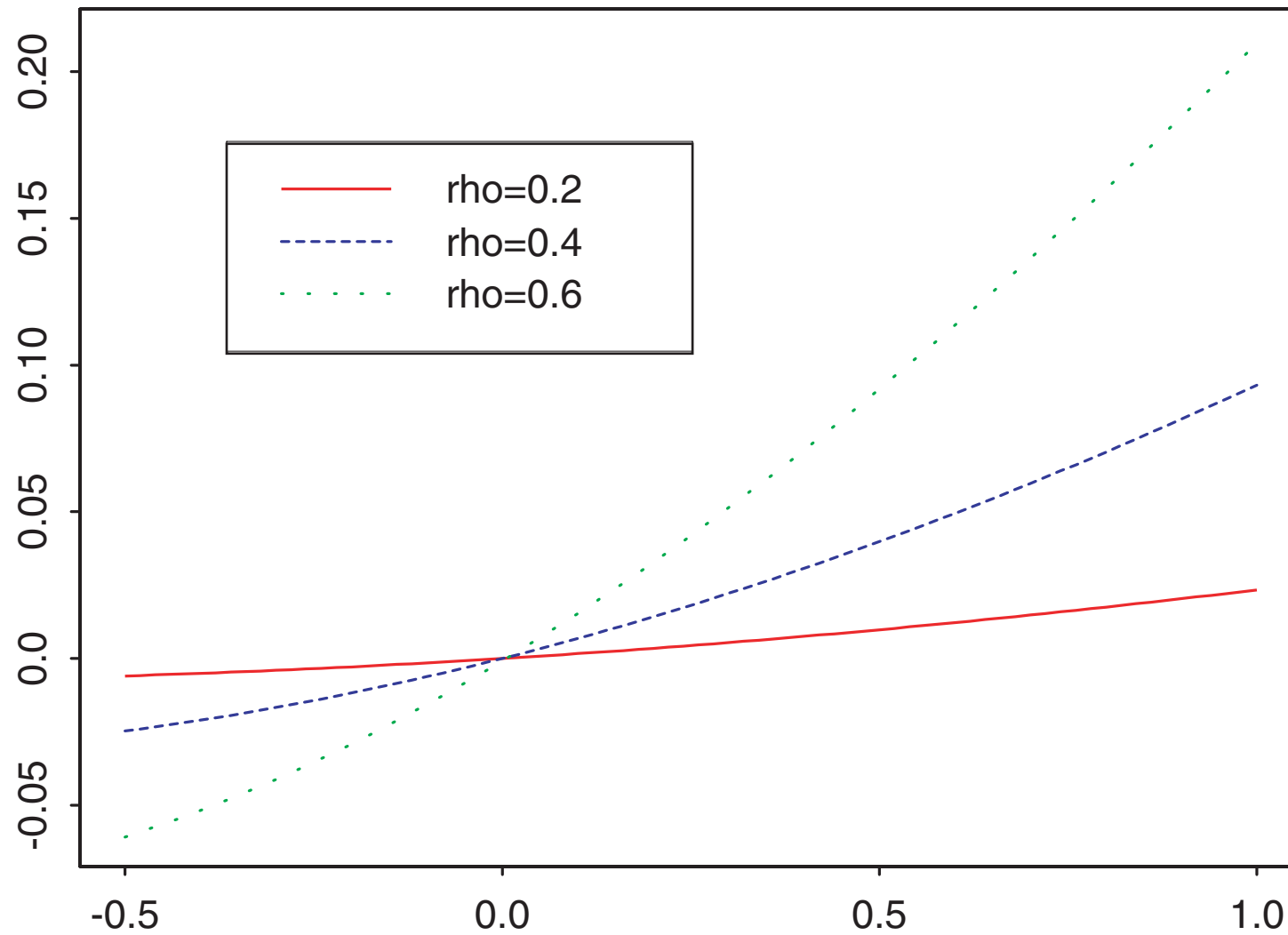
Influence on  $\sigma_V$  can be modeled by

$$(\sigma_V^{\text{new}})^2 = \sigma_{22}^2 + (1 + \Delta\sigma_X)^2 \sigma_{12}^2.$$

Relative difference  $\Delta\sigma_V$  in asset volatility:

$$\begin{aligned}\Delta\sigma_V &= \frac{\sigma_V^{\text{new}}}{\sigma_V^{\text{old}}} - 1 \\ &= (\sigma_V^{\text{old}})^{-1} \sqrt{\sigma_{22}^2 + (1 + \Delta\sigma_X)^2 \sigma_{12}^2} - 1.\end{aligned}$$

Relative difference in asset volatility as function of relative difference in foreign exchange rate volatility.



### 3 Comparison of different portfolio structures

Alternative factorization of the driving Brownian motion  $W$  in (1).

Special case of the correlation structure in (6):

Assume  $0 < \rho_* < \rho_1, \dots, \rho_K \leq 1$ .

Hence  $W_t \sim \mathcal{N}(0, R(n))$  with correlation matrix  $R(n) = \rho_{ij}(n)$

$$\rho_{ij}(n) = \begin{cases} 1, & \text{if } i = j, \\ \rho_h, & \text{if } i \text{ and } j \text{ are both in sector } h, \\ \rho_*, & \text{if } i \text{ and } j \text{ are in different sectors.} \end{cases}$$

$\rho_h$  gives the default correlation of counter-parties in sector  $h$  and  $\rho_*$  is the default correlation which is due to the global economic situation.

$\rho_h - \rho_*$  is the additional correlation in sector  $h$ .

Suppose that  $\rho_1 = \dots = \rho_K$ .

Aim: Comparison of two different portfolio structures  $s$  and  $s'$

$$s = (s_1, \dots, s_r, 0, \dots, 0), \quad s' = (s'_1, \dots, s'_l, 0, \dots, 0), \quad 1 \leq r, l \leq n$$

$r, l$ : number of sectors in these two portfolio structures,

$s_i, s'_i \in \mathbb{N}$ : number of firms in the two structures for all  $i$ ,

$$\sum_{i=1}^n s_i = \sum_{i=1}^n s'_i = n.$$

### Definition of a ordering of portfolio structures:

Denote by  $s_{[1]} \geq \dots \geq s_{[n]}$  the decreasing rearrangement of  $s$ , analogously for  $s'$ .  $s'$  majorizes  $s$  ( $s \prec s'$ ) if and only if

$$\sum_{i=1}^m s_{[i]} \leq \sum_{i=1}^m s'_{[i]}, \quad m = 1, \dots, n-1, \quad \text{and} \quad \sum_{i=1}^n s_{[i]} = \sum_{i=1}^n s'_{[i]}.$$



Theorem from Shaked, Tong (1992).

If  $s \prec s'$ , then  $E[L - c]^+ \leq E[L' - c]^+$  for all thresholds  $c \geq 0$ ,  
where  $L$  and  $L'$  describe the total loss of the different portfolios.  
(analogous to (9))

This means that well-balanced portfolios are less risky in the sense that the *expected loss shortfall* over a threshold  $c$  ( $c$  arbitrary) is smaller.

## Example\* (Effect of dependencies)

Suppose that the individual losses  $Y_i$  have a two-point distribution on 0 and 4, where the value 4 occurs with probability 0.06. The portfolio consists of  $n = 20$  risks.

scenario $m$	$s^{(m)}$
1	(1, 1, 1, ..., 1, 1, 1)
2	(4, 3, 3, 2, 2, 1, 1, 1, 1, 1, 1)
3	(8, 2, 2, 2, 2, 2, 2)
4	(4, 4, 4, 3, 3, 2)
5	(15, 2, 1, 1, 1)
6	(5, 5, 5, 5)
7	(10, 5, 5)
8	(20)

Table 1

\*See Bäuerle/Müller (1998)

Relative expected loss excess (divided by the independent case  $m = 1$ ) multiplied by 100 for several thresholds for the 8 different scenarios in Table 1:

$$(100E[\sum_{i=1}^{20} Y_i^{(m)} - c]^+ / E[\sum_{i=1}^{20} Y_i^{(1)} - c]^+).$$

Note that the expectation of the total loss equals 4.8 and the outcomes range between 0 and 80.

threshold	scenario							
	$s^{(1)}$	$s^{(2)}$	$s^{(3)}$	$s^{(4)}$	$s^{(5)}$	$s^{(6)}$	$s^{(7)}$	$s^{(8)}$
0	100	100	100	100	100	100	100	100
1	100	105	109	110	111	112	113	116
2	100	113	121	124	126	129	132	139
3	100	124	140	145	150	155	161	173
4	100	144	173	182	191	200	210	233
6	100	174	210	229	272	272	295	347
8	100	270	330	385	537	506	572	717
10	100	327	478	480	830	700	834	1128

Table 2

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